# Estimating Latent Variable Graphical Models with Moments and Likelihoods

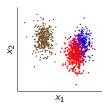
Arun Tejasvi Chaganty Percy Liang

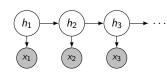
Stanford University

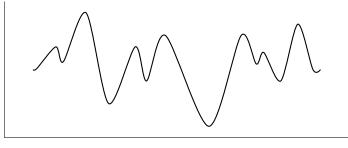
June 22, 2014

## Latent Variable Graphical Models

- Gaussian Mixture Models
- ► Latent Dirichlet Allocation
- Hidden Markov Models
- PCFGs
- **.** . . .

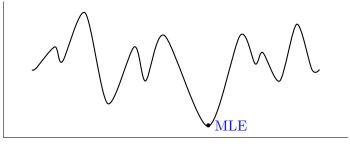






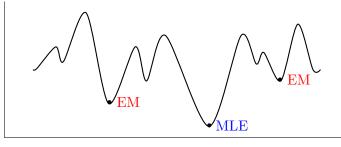
parameters

▶ Log-likelihood function is non-convex.



#### parameters

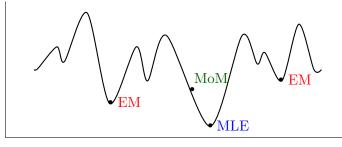
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- MLE is consistent but intractable.



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- ► Local methods (EM, gradient descent, ...) are tractable but inconsistent.

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- Log-likelihood function is non-convex.
- MLE is consistent but intractable.
- Local methods (EM, gradient descent, ...) are tractable but inconsistent.
- Method of moments estimators can be consistent and computationally-efficient, but require more data.

- Several estimators based on the method of moments.
  - Phylogenetic trees: Mossel and Roch 2005.
    - ▶ Hidden Markov models: Hsu, Kakade, and Zhang 2009
  - Latent trees: Anandkumar et al. 2011
  - ▶ Latent Dirichlet Allocation: Anandkumar et al. 2012
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- ▶ These estimators are applicable only to a specific type of model.

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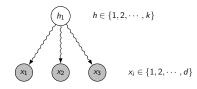
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- ▶ These estimators are applicable only to a specific type of model.
- In contrast, EM and gradient descent apply for almost any model.
- Note: some work in the observable operator framework does apply to a more general model class.
  - ▶ Weighted automata: Balle and Mohri 2012.
  - Junction trees: Song, Xing, and Parikh 2011
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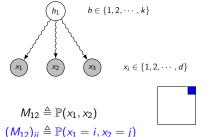


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- ▶ How can we apply the method of moments to estimate parameters efficiently for a general model?

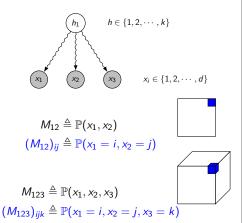
▶ Discrete models with k hidden and  $d \ge k$  observed values.



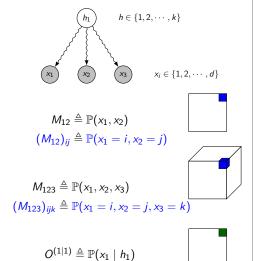
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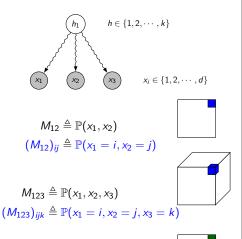


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 $(O^{(1|1)})_{ii} \triangleq \mathbb{P}(x_1 = i \mid h_1 = i)$ 

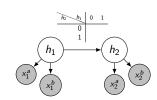
- ▶ Discrete models with k hidden and d > k observed values.
- Parameters and marginals can be represented as matrices and tensors.
- Presented in terms of infinite data and exact moments.



$$O^{(1|1)} \triangleq \mathbb{P}(x_1 \mid h_1)$$
  
 $(O^{(1|1)})_{ij} \triangleq \mathbb{P}(x_1 = i \mid h_1 = j)$ 

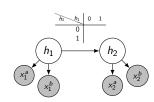


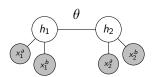
 Directed models parameterized by conditional probability tables.



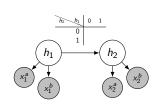
 $\theta$ 

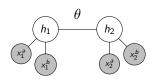
- Directed models parameterized by conditional probability tables.
- Undirected models parameterized as a log-linear model. Identify modulo  $A(\theta)$ .





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- Focus on directed models, but return to undirected models later.

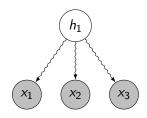




## Background: Three-view mixture models aka bottlenecks

#### Definition (Bottleneck)

A hidden variable h is a **bottleneck** if there exist three observed variables (**views**)  $x_1, x_2, x_3$  that are conditionally independent given h.

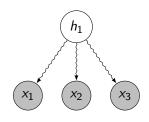


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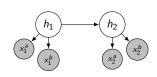
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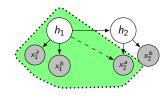
- Anandkumar et al. 2013 provide an algorithm to estimate conditional moments  $O^{(i|1)} \triangleq \mathbb{P}(x_i \mid h_1)$  based on tensor eigendecomposition.
- In general, three views are necessary for identifiability (Kruskal 1977).



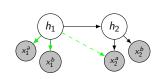
Each edge has a set of parameters.



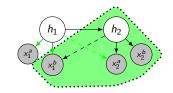
- ► Each edge has a set of parameters.
- ▶  $h_1$  and  $h_2$  are bottlenecks.



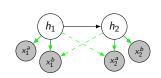
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- $\blacktriangleright$   $h_1$  and  $h_2$  are bottlenecks.
- We can learn  $\mathbb{P}(x_1^a|h_1), \mathbb{P}(x_1^b|h_1), \cdots$



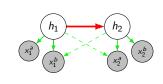
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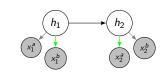
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- We can learn  $\mathbb{P}(x_1^a|h_1), \mathbb{P}(x_1^b|h_1), \cdots$
- ► However, we can't learn  $\mathbb{P}(h_2|h_1)$  this way.

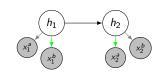


► Observe the joint distribution of  $x_1, x_2$ ,



$$\underbrace{\mathbb{P}(x_1^b, x_2^a)}_{M_{12}} = \sum_{h_1, h_2} \underbrace{\mathbb{P}(x_1^b \mid h_1) \mathbb{P}(x_2^a \mid h_2) \mathbb{P}(h_1, h_2)}_{O^{(2|2)}}.$$

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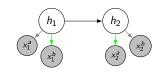
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$$M_{12} = O^{(1|1)}$$



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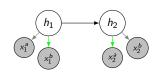








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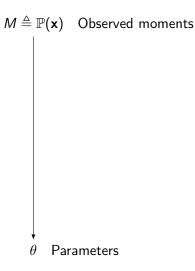
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- ▶ Normalize for  $\mathbb{P}(h_2 \mid h_1)$ .



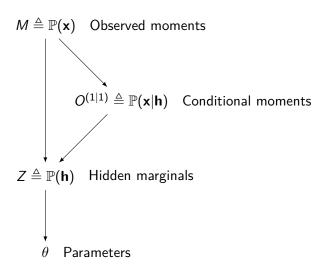


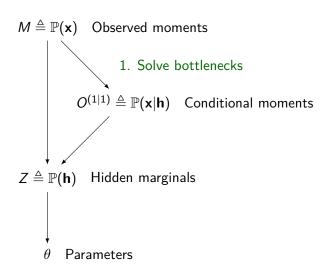


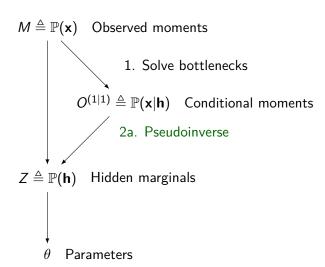










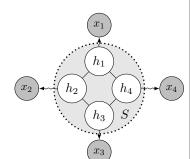


#### **Exclusive Views**

## Definition (Exclusive views)

We say  $h_i \in S \subseteq \mathbf{h}$  has an **exclusive view**  $x_v$  if

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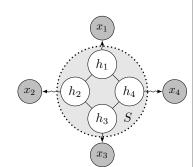


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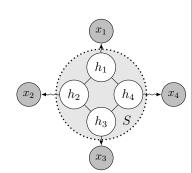


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- 3. A set has exclusive views if each  $h_i \in S$  has an exclusive view.



#### Exclusive views give parameters

▶ Given exclusive views,  $\mathbb{P}(x \mid h)$ , learning cliques is solving a linear equation!

$$\underbrace{\mathbb{P}(x_1,\ldots,x_m)}_{M} = \sum_{h_1,\ldots,h_m} \underbrace{P(x_1|h_1)}_{O^{(1|1)}} \cdots \underbrace{P(h_1,\cdots,h_m)}_{Z}$$

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$$\underbrace{O^{(3|3)}}_{M}$$

$$= \underbrace{O^{(2|2)}}_{O^{(2|2)}}$$

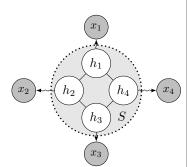


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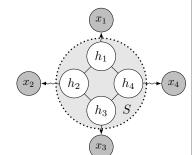
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\underbrace{O^{(3|3)}}_{D^{(2|2)}} \rightarrow \underbrace{Z} = \underbrace{O^{(1|1)\dagger}}_{M} \qquad \underbrace{O^{(2|2)\dagger}}_{D^{(2|2)}}$$

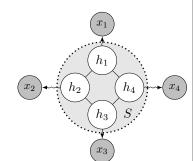
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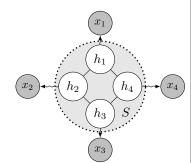
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- Theorem: A clique in which each hidden variable is a bottleneck has exclusive views.

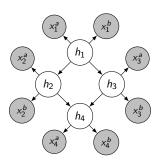


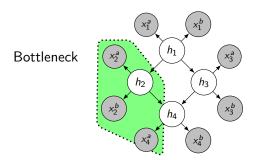
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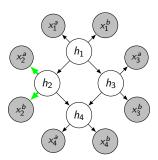


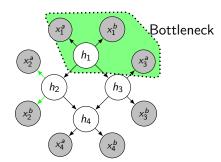
- When are we assured exclusive views?
- Theorem: A clique in which each hidden variable is a bottleneck has exclusive views.
  - Follows by graph independence conditions.
  - We say that the clique is "bottlenecked".

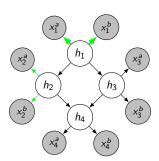


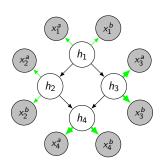


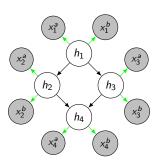


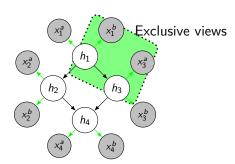


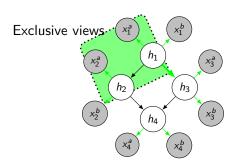


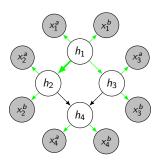


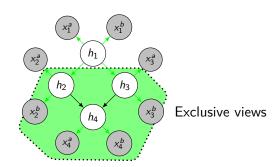


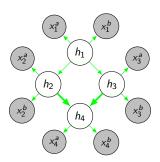


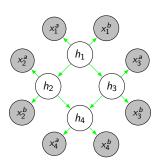






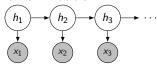




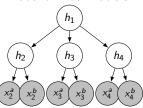


# More Bottlenecked Examples

#### Hidden Markov models

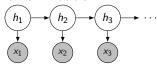


#### Latent Tree models

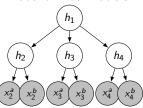


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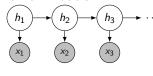


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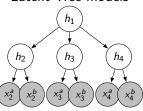


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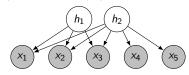
#### Hidden Markov models



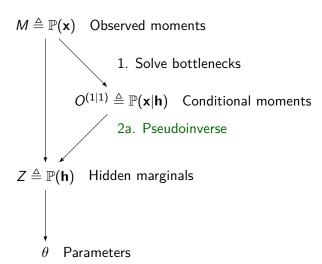
#### Latent Tree models



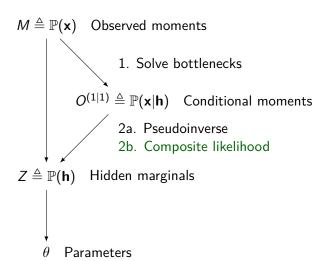
#### Noisy Or (non-example) (Halpern and Sontag 2013)



#### Outline

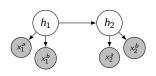


#### Outline



# Convex marginal likelihoods

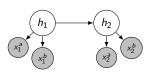
The MLE is statistically most efficient, but usually non-convex.

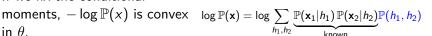


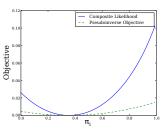
$$\log \mathbb{P}(\mathbf{x}) = \log \sum_{h_1,h_2} \mathbb{P}(\mathbf{x}_1|h_1) \mathbb{P}(\mathbf{x}_2|h_2) \mathbb{P}(h_1,h_2)$$

# Convex marginal likelihoods

- ▶ The MLE is statistically most efficient, but usually non-convex.
- If we fix the conditional in  $\theta$ .

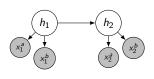




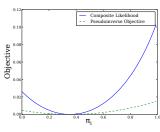


# Convex marginal likelihoods

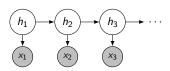
- ▶ The MLE is statistically most efficient, but usually non-convex.
- If we fix the conditional in  $\theta$ .
- No closed form solution, but a local method like EM is guaranteed to converge to the global optimum.



moments, 
$$-\log \mathbb{P}(x)$$
 is convex  $\log \mathbb{P}(x) = \log \sum_{h_1,h_2} \underbrace{\mathbb{P}(x_1|h_1)\mathbb{P}(x_2|h_2)}_{\text{known}} \mathbb{P}(h_1,h_2)$ 

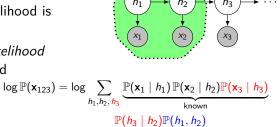


► In general, the full likelihood is still non-convex.

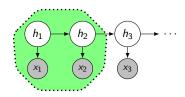


$$\log \mathbb{P}(\mathbf{x}_{123}) = \log \sum_{h_1, h_2, h_3} \underbrace{\mathbb{P}(\mathbf{x}_1 \mid h_1) \mathbb{P}(\mathbf{x}_2 \mid h_2) \mathbb{P}(\mathbf{x}_3 \mid h_3)}_{\text{known}}$$
$$\mathbb{P}(h_3 \mid h_2) \mathbb{P}(h_1, h_2)$$

- ► In general, the full likelihood is still non-convex.
- ► Consider composite likelihood on a subset of observed variables.  $\log \mathbb{P}(\mathbf{x}_{12})$

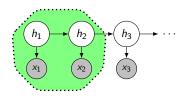


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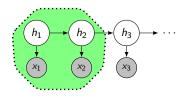
$$\log \mathbb{P}(\mathbf{x}_{12}) = \log \sum_{h_1,h_2} \underbrace{\mathbb{P}(\mathbf{x}_1 \mid h_1) \mathbb{P}(\mathbf{x}_2 \mid h_2)}_{\text{known}}$$
$$\mathbb{P}(h_1,h_2)$$

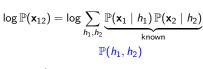
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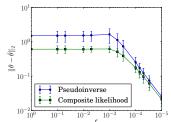


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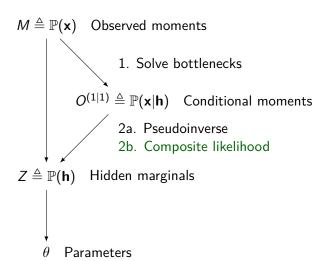
- ► In general, the full likelihood is still non-convex.
- Consider composite likelihood on a subset of observed variables.
- Can be shown that estimation with composite likelihoods is consistent (Lindsay 1988).
- Asymptotically, the composite likelihood estimator is more efficient.



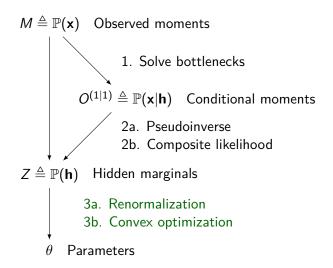




#### Outline



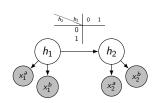
## Outline



# Recovering parameters in directed models

- Conditional probability tables are the default for a directed model.
- Can be recovered by normalization:

$$\mathbb{P}(h_2 \mid h_1) = \frac{\mathbb{P}(h_1, h_2)}{\sum_{h_2} \mathbb{P}(h_1, h_2)}.$$

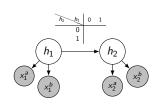


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$$\mathbb{P}(h_2 \mid h_1) = \frac{\mathbb{P}(h_1, h_2)}{\sum_{h_2} \mathbb{P}(h_1, h_2)}.$$

No dependence on tree-width. Memory, computation and samples depend linearly on the size of each clique.



Assume a log-linear parameterization,

$$p_{ heta}(\mathbf{x}, \mathbf{h}) = \exp\left(\sum_{\mathcal{C} \in \mathcal{G}} \mathbf{\theta}^{ op} \phi(\mathbf{x}_{\mathcal{C}}, \mathbf{h}_{\mathcal{C}}) - A(\mathbf{\theta})\right).$$

Assume a log-linear parameterization,

$$p_{ heta}(\mathbf{x}, \mathbf{h}) = \exp\left(\sum_{\mathcal{C} \in \mathcal{G}} \theta^{\top} \phi(\mathbf{x}_{\mathcal{C}}, \mathbf{h}_{\mathcal{C}}) - A(\theta)\right).$$

The *unsupervised* negative log-likelihood is non-convex,

$$\mathcal{L}_{\mathsf{unsup}}(\theta) \triangleq \mathbb{E}_{\mathbf{x} \sim \mathcal{D}}[-\log \sum_{\mathbf{h} \in \mathcal{H}} p_{\theta}(\mathbf{x}, \mathbf{h})].$$

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▶ However, the *supervised* negative log-likelihood is convex,

$$\begin{split} \mathcal{L}_{\mathsf{sup}}(\theta) &\triangleq \mathbb{E}_{(\mathbf{x}, \mathbf{h}) \sim \mathcal{D}_{\mathsf{sup}}} \left[ -\log p_{\theta}(\mathbf{x}, \mathbf{h}) \right] \\ &= -\theta^{\top} \left( \sum_{\mathcal{C} \in \mathcal{G}} \mathbb{E}_{(\mathbf{x}, \mathbf{h}) \sim \mathcal{D}_{\mathsf{sup}}} [\phi(\mathbf{x}_{\mathcal{C}}, \mathbf{h}_{\mathcal{C}})] \right) + A(\theta). \end{split}$$

Recall, the marginals can typically estimated from supervised data.

$$\mathcal{L}_{\mathsf{sup}}(\theta) = -\theta^{\top} \underbrace{\left( \sum_{\mathcal{C} \in \mathcal{G}} \mathbb{E}_{(\mathbf{x},\mathbf{h}) \sim \mathcal{D}_{\mathsf{sup}}} [\phi(\mathbf{x}_{\mathcal{C}},\mathbf{h}_{\mathcal{C}})] \right)}_{\mu_{\mathcal{C}}} + \textit{A}(\theta).$$

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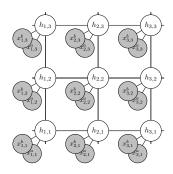
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However, the marginals can also be *consistently* estimated by moments!

$$\mu_{\mathcal{C}} = \sum_{\mathbf{x}_{\mathcal{C}}, \mathbf{h}_{\mathcal{C}}} \underbrace{\mathbb{P}(\mathbf{x}_{\mathcal{C}} \mid \mathbf{h}_{\mathcal{C}})}_{\text{moments hidden marginals}} \phi(\mathbf{x}_{\mathcal{C}}, \mathbf{h}_{\mathcal{C}}).$$

# Optimizing pseudolikelihood

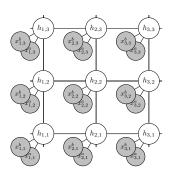
▶ Estimating  $\mu_C$ : independent of treewidth.



# Optimizing pseudolikelihood

- ▶ Estimating  $\mu_C$ : independent of treewidth.
- ► Computing  $A(\theta)$ : dependent on treewidth.

$$A(\theta) \triangleq \log \sum_{\mathbf{x}, \mathbf{h}} \exp \left( \theta^{\top} \phi(\mathbf{x}, \mathbf{h}) \right).$$



# Optimizing pseudolikelihood

- Estimating  $\mu_{\mathcal{C}}$ : independent of treewidth.
- ► Computing  $A(\theta)$ : dependent on treewidth.

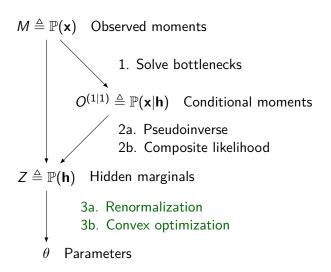
$$A(\theta) \triangleq \log \sum_{\mathbf{x}, \mathbf{h}} \exp \left( \theta^{\top} \phi(\mathbf{x}, \mathbf{h}) \right).$$

Instead, use pseudolikelihood (Besag 1975) to consistently estimate distributions over local neighborhoods.

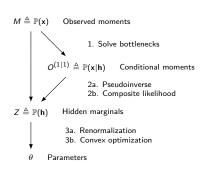
$$\begin{pmatrix} h_{1,3} & h_{2,3} & h_{3,3} \\ x_{1,3}^b & x_{2,3}^b & x_{3,2}^b \\ h_{1,2} & h_{2,2} & h_{3,2} \\ x_{1,2}^b & x_{2,2}^b & x_{3,2}^b \\ h_{1,1} & x_{2,1}^b & x_{3,2}^b \\ x_{1,1}^b & x_{2,1}^b & x_{3,1}^b \\ x_{1,1}^b & x_{2,1}^b & x_{3,2}^b \\ x_{1,2}^b & x_{2,2}^b & x_{2,2}^b \\ x_{1$$

$$A_{\mathsf{pseudo}}(\theta; \mathcal{N}(\textit{a})) \triangleq \log \sum \exp \left( \theta^\top \phi(\mathbf{x}_{\mathcal{N}}, \mathbf{h}_{\mathcal{N}}) \right).$$

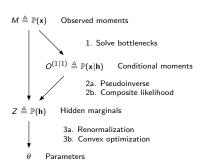
## Outline



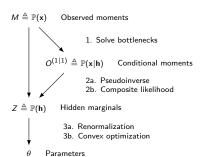
An algorithm for any bottlenecked discrete graphical model.



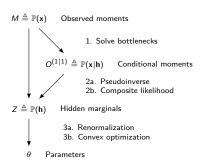
- An algorithm for any bottlenecked discrete graphical model.
- Combine moment estimators with likelihood estimators.



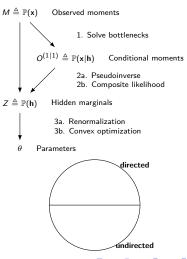
- An algorithm for any bottlenecked discrete graphical model.
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- Extends to log-linear models.



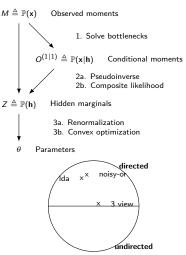
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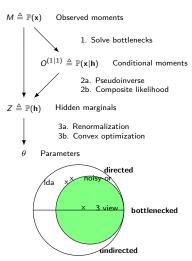
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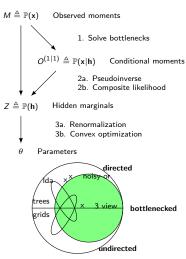
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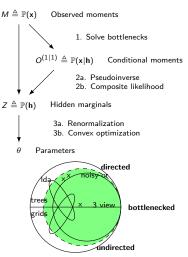
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- Thank you! Poster: M58

